

1.

Suppose, we have Y_1, Y_2, \dots, Y_T i.i.d. random variables (r.v) with mean β and variance σ_y^2 .

THEORY:

$$\text{mean of } Y_t : E(Y_t) = \beta$$

$$\text{variance of } Y_t : E(Y_t - E(Y_t))^2 = \sigma_y^2$$

PRACTICE = ESTIMATION:

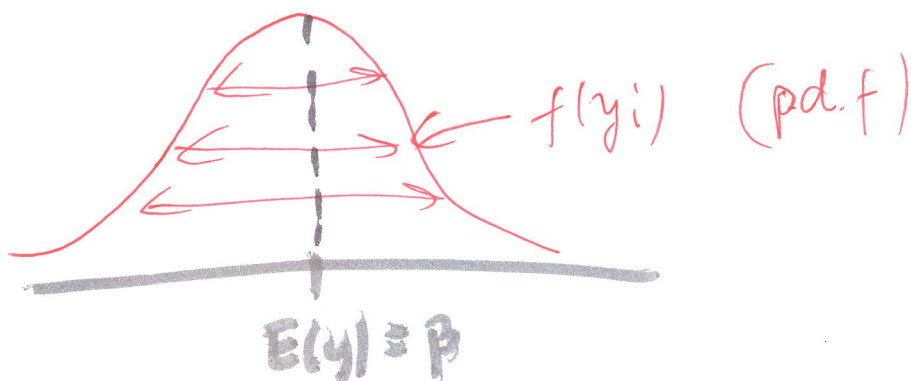
Suppose we have a random sample of observations of Y_i : y_1, y_2, \dots, y_T

• mean estimator

$$\bar{y} = \hat{\beta} = \frac{1}{T} \sum_{t=1}^T y_t$$

• variance estimator

$$\hat{\sigma}_y^2 = \frac{1}{T} \sum_{t=1}^T (y_t - \hat{\beta})^2$$



get the data: y_1, y_2, \dots, y_T

② use OLS to estimate β

$$y_t = \beta + \underline{e_t}, \quad e_t \sim (0, \sigma_e^2)$$

$$\text{OLS: } \min \sum_{t=1}^T e_t^2 = \min \sum_{t=1}^T (y_t - \beta)^2$$

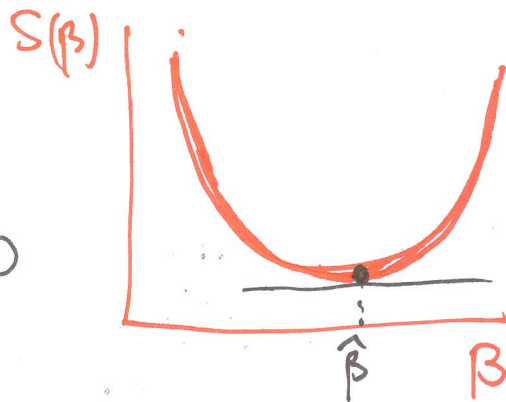
$$\min S(\beta) = \sum_{t=1}^T (y_t - \beta)^2$$

$$= \sum_{t=1}^T y_t^2 - 2\beta \sum_{t=1}^T y_t + T\beta^2 \quad (a-b)^2$$

$\Rightarrow \sum_{t=1}^T \beta^2 = T\beta^2$

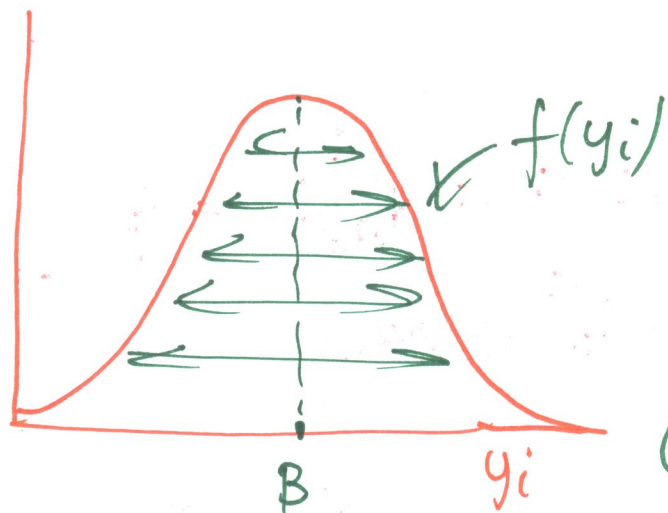
$$\frac{dS(\beta)}{d\beta} = -2 \sum_{t=1}^T y_t + 2T\beta$$

$$\frac{dS(\beta)}{d\beta} = 0 \Leftrightarrow -2 \sum y_t + 2T\beta = 0$$



replace β by $\hat{\beta}$, $b = ?$

$$\hat{\beta} = \frac{\sum y_t}{T} = \frac{1}{T} \sum y_t = \bar{y}$$



$$\beta = \bar{y}$$

$$\sigma_y^2$$
$$\hat{\sigma}_y^2$$

Properties of the estimator:

3.

- $E(\hat{\beta})$ we want $E(\hat{\beta}) = \beta$
if yes, $\hat{\beta}$ is UNBIASED

$$E(\hat{\beta}) = E\left(\frac{\sum y_t}{T}\right) = \frac{1}{T} \sum_{t=1}^T E y_t = \frac{1}{T} T \beta = \beta$$

$$\begin{aligned} \bullet \text{var}(\hat{\beta}) &= \text{var}\left(\frac{1}{T} \sum y_t\right) = \left(\frac{1}{T}\right)^2 \text{var} \sum y_t \\ &= \frac{1}{T^2} \sum_{t=1}^T \text{var} y_t \end{aligned}$$

$$\text{var}(\hat{\beta}) = \frac{1}{T^2} T \cdot \sigma_y^2 = \frac{\sigma_y^2}{T}$$

efficient estimator = minimum variance estimator.

OLS \rightarrow yes!

in this model only! $\sigma_e^2 = \sigma_y^2$. It estimates the MARGINAL mean and variance of y_t .

An alternative model for y_t :

Suppose: y_t : spending
 x_t : income

$$y_t = \beta_1 + \beta_2 \underline{x_t} + e_t$$

ASSUMPTIONS ON e_t :

$$e_t \sim (0, \sigma_e^2)$$

$$E(e_t | x_t) = 0$$

allows us to estimate the **CONDITIONAL = MODEL**
BASED MOMENTS OF y_t :

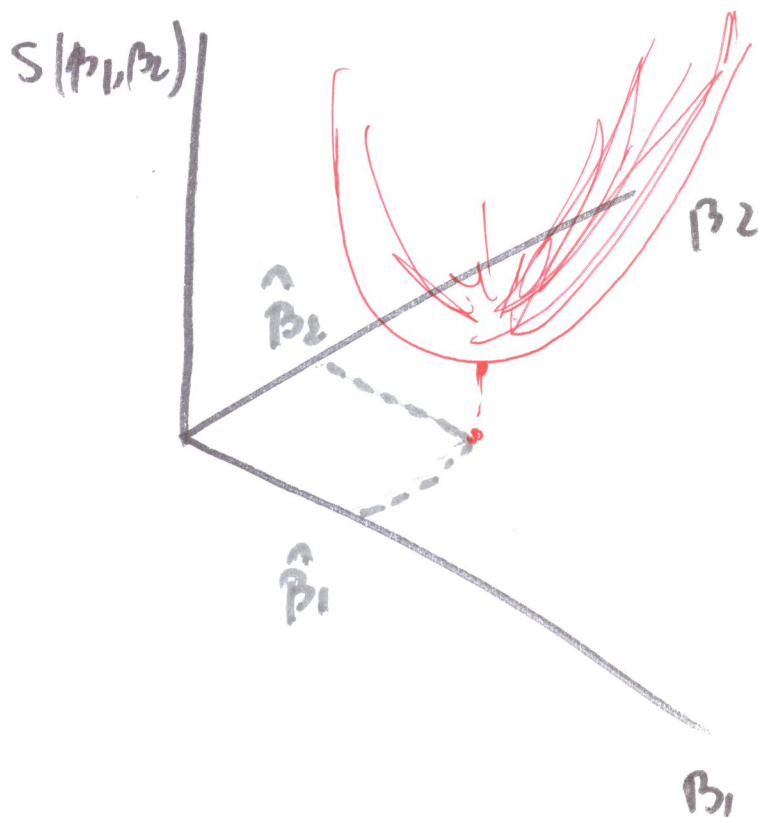
$$E(y_t | x_t) = \beta_1 + \beta_2 x_t$$

$$\begin{aligned} \text{var}(y_t | x_t) &= E \left(y_t - \underline{(\beta_1 + \beta_2 x_t)} \right)^2 = E(e_t^2) = \sigma_e^2 \\ &= E \left[\left(y_t - \underline{E(y_t | x_t)} \right) | x_t \right]^2 \end{aligned}$$

$$y_t = \beta_1 + \beta_2 x_t + e_t, \quad e_t \sim (0, \sigma e^2)$$

OLS:

$$\min_{\beta_1, \beta_2} S(\beta_1, \beta_2) = \min \sum_{t=1}^T e_t^2 = \min \sum_{t=1}^T (y_t - \beta_1 - \beta_2 x_t)^2$$



$$\hat{\beta}_1 = \bar{y} - \hat{\beta}_2 \bar{x}$$

$$\hat{\beta}_2 = \frac{\widehat{\text{cov}}(y_t, x_t)}{\widehat{\text{var}} x_t} = \frac{\frac{1}{T} \sum (y_t - \bar{y})(x_t - \bar{x})}{\frac{1}{T} \sum (x_t - \bar{x})^2}$$

properties of OLS

$$E(\hat{\beta}_1) = \beta_1$$

$$E(\hat{\beta}_2) = \beta_2$$

~~data~~ unbiased

Greene. ✓

$$E(\hat{\beta}_1 | x_t)$$

$$E(\hat{\beta}_2 | x_t)$$

$$y_t = \beta_1 + \epsilon_t$$

$$y_t = \beta_1 + \beta_2 x_t + \epsilon_t$$